

# Early health technology assessment of future clinical decision rule aided triage of patients presenting with acute chest pain in primary care

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# Early health technology assessment of future clinical decision rule aided triage of patients presenting with acute chest pain in primary care

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The objective of the paper is to estimate the number of patients presenting with chest pain suspected of acute coronary syndrome (ACS) in primary care and to calculate possible cost effects of a future clinical decision rule (CDR) incorporating a point-of-care test (PoCT) as compared with current practice. The annual incidence of chest pain, referrals and ACS in primary care was estimated based on a literature review and on a Dutch and Belgian registration study. A health economic model was developed to calculate the potential impact of a future CDR on costs and effects (ie, correct referral decisions), in several scenarios with varying correct referral decisions. One-way, two-way, and probabilistic sensitivity analyses were performed to test robustness of the model outcome to changes in input parameters. Annually, over one million patient contacts in primary care in the Netherlands concern chest pain. Currently, referral of eventual ACS negative patients (false positives, FPs) is estimated to cost €1,448 per FP patient, with total annual cost exceeding 165 million Euros in the Netherlands. Based on 'international data', at least a 29% reduction in FPs is required for the addition of a PoCT as part of a CDR to become cost-saving, and an additional €16 per chest pain patient (ie, 16.4 million Euros annually in the Netherlands) is saved for every further 10% relative decrease in FPs. Sensitivity analyses revealed that the model outcome was robust to changes in model inputs, with costs outcomes mainly driven by costs of FPs and costs of PoCT. If PoCT-aided triage of patients with chest pain in primary care could improve exclusion of ACS, this CDR could lead to a considerable reduction in annual healthcare costs as compared with current practice.

**Key words:** biomarkers; cardiovascular disease; clinical decision rule; cost effects; emergency medicine; primary care

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## Background

A clinical decision rule (CDR) based on history and physical examination to safely rule out acute coronary syndrome (ACS) in primary care is not

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available (Grijseels *et al.*, 1996; Winters and Katzen, 2006; Body *et al.*, 2010; Gencer *et al.*, 2010; Bosner *et al.*, 2013). Therefore, in primary care, up to 73% of patients with new or altered chest pain are immediately referred by the GP to the emergency department (Bruins Slot *et al.*, 2013). However, only a minority of those patients (up to 26% in literature) are subsequently diagnosed with an acute life threatening disease, for example ACS ('true referrals' or 'true positives' (TPs)) (Chambers and Bass, 1998; Buntinx *et al.*, 2001; Kohn *et al.*, 2005; Hocaoglu *et al.*, 2008; Bruins Slot *et al.*, 2013). Patients that were referred and were found to be ACS negative ('false referrals' or 'false positives' (FPs)) were diagnosed with alternative diseases with advantageous courses (Buntinx *et al.*, 2001; Kohn *et al.*, 2005). In this context, the term 'false' is used to indicate a referral in absence of ACS afterwards. The referral itself however is undisputed, as it is the result of a GP unable to exclude a potentially life threatening disease. On the other hand, incidentally, ACS is present in patients that were initially not referred ('false non-referrals' or 'false negatives' (FNs)) (Buntinx *et al.*, 2001; Kohn *et al.*, 2005).

ACS negative referrals (FPs) pose a significant burden on healthcare resources, and reduction of FP can lead to increased patient comfort, while decreasing costs (Graff *et al.*, 1997; Dumville *et al.*, 2007; Mourad *et al.*, 2013). Therefore, specificity of a future CDR for ACS should be higher – and sensitivity should at the least be maintained – as compared with current practice that is based on a GPs' clinical judgement only. Recently, the efficiency of similar diagnostic processes in primary care was improved by introducing a cost-effective CDR, combining point-of-care tests (PoCTs) with clinical findings, leading to less prescription of unnecessary antibiotics in lower respiratory tract infections and less unnecessary referral for suspected pulmonary embolism (Cals *et al.*, 2011; Geersing *et al.*, 2012; Little *et al.*, 2013).

The availability of a validated CDR, incorporating a PoCT measuring a biomarker of myocardial damage (eg, high-sensitive troponin (hsTn) or heart-type fatty acid binding protein (H-FABP)) is anticipated, however not yet available (Nilsson *et al.*, 2013; Glatz and Mohren, 2013; Glatz and Renneberg, 2014; Willemsen *et al.*, 2014, 2015). The majority of GPs expect future PoCTs to be of added value in ruling out ACS (Kip *et al.*, 2017).

## Objectives

- (1) To estimate the number of patients with chest pain in primary care, their referral rates and the incidence of ACS among these patients
- (2) To assess the minimum required reduction in ACS negative referrals (FPs) due to a future CDR, incorporating clinical findings and a PoCT, to become cost-saving, assuming that the number of non-referrals among ACS patients (FNs) equals current practice
- (3) To assess the impact of a relative decrease in ACS negative referrals (FPs), with decrements of 10%
- (4) To assess the combined impact of simultaneously varying the referrals among non-ACS patients (FPs) and the costs of the PoCT, to determine which combinations are expected to save costs

## Methods

### Estimation of annual patient numbers based on literature

International data and a Dutch registration study were used as separate sources to estimate numbers of true and false referrals (referred to as true and false positives respectively (TPs and FPs)) as well as true and false non-referrals (referred to as true and false negatives respectively (TNs and FN)) in primary care in the Netherlands (Hoorweg *et al.*, 2017). The international data were obtained through an extensive literature search. Those data will be referred to as 'international data', and will be used as the base case scenario in the remainder of this paper. However, as registration studies are rare, it was expected that part of the data were to be derived from studies describing relevant data that were primarily designed to meet other objectives. Therefore, we searched PubMed and Embase from January 1989 to May 2017 for chest complaints in primary care (see Supplementary Material Figure S1 for an overview of the literature search and strategy). Articles assumed relevant based on title/abstracts were read in order to select all studies supplying relevant data on the incidence of chest complaints in primary care, referral rates, and final diagnoses (when available). Additional relevant articles were identified from the references in these selected

papers. Besides these sources, we used Dutch registry data (governmental data, data from the Dutch Heart Foundation and data from the Dutch Central Statistics Agency) (Okkes *et al.*, 2005; Academic Hospital Maastricht, 2013; Catholic University Leuven, 2013; Dutch Heart Foundation, 2013; Centraal Bureau voor de Statistiek (CBS), 2014a; 2014b; 2014c; 2014d). Data from the literature review as well as from the Dutch registries were pooled and mean values with 95% confidence intervals (95% CI) were calculated. In most cases, the included studies did not describe all probabilities in the patient pathway (eg, the set of studies that described the percentage of patients that is referred to secondary care, differed from the set of studies that described the proportion of ACS positive cases among those referred patients). Therefore, the sample sizes of the pooled estimates differed across the different parameters. Eventually, the pooled international data were translated to estimated referral rates and incidence of ACS in the Netherlands. As a second scenario to model the patient pathway of Dutch patients with chest pain in primary care, a Dutch/Belgian registration study was used, and this scenario will be referred to as 'NL and B data' (Hoorweg *et al.*, 2017).

### Model diagnostic process, resource use

A health economic model was developed to estimate the costs of the full diagnostic work-up. Costs were estimated based on several sources (see Table 1), and expressed in 2016 Euros (CBS, 2016). The analysis was performed from a health-care perspective, incorporating all direct medical

costs that occur from the moment a patient presents with chest pain in primary care, until the patient was either referred to and diagnosed and treated in secondary care, or sent home following the GP consultation (without referral). As this time horizon is less than one year, discounting of costs and effects was not required. Calculations were performed using unit costs for assessment in primary care including PoCT testing, ambulance transport to the hospital, and assessment in secondary care. Exclusion of ACS by a GP costs €18 without PoCT, and increases to €63 when cost for usage of a PoCT of €45 is included (based on the expected price of a H-FABP PoCT which is currently in development). The cost for every patient that is assessed in a hospital for cardiac analysis with and without the eventual presence of an underlying ACS is estimated at €5735 and €1426 respectively, including hospital transport by ambulance.

### Outcome measures

The primary effectiveness measure was defined as the percentage of patients in whom ACS (including both unstable angina (UA) and acute myocardial infarction (AMI)) was correctly diagnosed or excluded when using the CDR as compared with current practice. The incremental cost-effectiveness ratio (ICER) was therefore expressed as incremental costs per patient when using a CDR (including PoCT), as compared with current practice, and divided by the difference in the number of patients in whom the correct referral decision is made in both work-ups.

**Table 1** Model input: cost data

Parameter	Value [95% CI]	Distribution
Consultation at tariff GP (double)*	€18.00 [€10.18–27.93]	Gamma
PoCT (including finger prick needle, and VAT)**	€45.00 [€26.08–69.52]	Gamma
Ambulance transport medium to high urgency (medical personnel A1/A2 drive, overhead costs call center, and VAT)***	€750.00 [€429.70–1,172.51]	Gamma
Analysis CCU, no ACS (diagnostic tests, medical personnel, hospital stay one to two days)****	€676.00 [€386.75–1054.49]	Gamma
Analysis CCU, ACS present (diagnostic tests, medical personnel, hospital stay 3 days, PCI)****	€4985.00 [€2823.36–7673.56]	Gamma

Cost prices for relevant events in in- or excluding ACS in The Netherlands. Cost prices are based on the following sources: double consultation price GP in The Netherlands (\*), estimations from manufacturer (\*\*), cost price requested at large ambulance service in South of The Netherlands (\*\*\*), average diagnosis-treatment combination tariffs of considerable number of Dutch hospitals of varying type (small, large, academic, urban, rural) (\*\*\*\*).

95% CI = 95% confidence interval; ACS = acute coronary syndrome; CCU = coronary care unit; H-FABP = heart-type fatty acid binding protein; PCI = percutaneous coronary intervention; PoCT = point-of-care test; VAT = value-added tax.

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### One- and two-way sensitivity analysis

A two-way deterministic sensitivity analysis was performed to obtain insight into the combined impact of simultaneously varying the cost of the PoCT and the %FPs (Briggs *et al.*, 2006). For subsequent analysis (ie, one-way and probabilistic sensitivity analyses), the minimum required relative decrease in FPs to obtain a strategy that is cost-saving compared with current practice was applied in the scenarios with PoCT (while assuming that costs of the PoCT remain unaffected). Following this, to identify which individual cost parameters drive the model outcome, given fixed costs of the PoCT and the minimum required relative decrease in FPs, we conducted a one-way deterministic sensitivity analysis (Briggs *et al.*, 2006). In the one-way sensitivity analysis, the impact of a 25% decrease and increase in all cost input parameters on the costs per patient presenting with suspected ACS in primary care was analysed. As the impact on FP referrals in the CDR + PoCT strategy was arbitrarily chosen (based on the minimum required reduction in FPs), it was decided to only incorporate the impact on costs in

this one-way sensitivity analysis. In addition, this avoids double counting, as both the numerator and the denominator of the ICER are affected by a change in correct referral decisions.

### Probabilistic sensitivity analysis

Distributions were assigned to all model parameters (Briggs *et al.*, 2006). Subsequently, random samples were drawn for all model input parameters simultaneously. An overview of the type of distribution used for each input parameter, as well as the accompanying 95% confidence intervals (95% CI) is provided in Tables 1 and 2. A probabilistic sensitivity analysis (PSA) based on Monte Carlo simulation with 10 000 samples was performed to determine the effect of joint uncertainty in all model input parameters on model outcome.

## Results

### Estimation of relevant patient numbers

In our literature search, 1900 articles were assessed. The majority of articles were not

**Table 2** Model input: effectiveness data for three different base cases

Scenario (estimated annual number of chest pain patients for the Netherlands [95% CI])	Referred		Not referred		Distribution
	ACS (TP) [95% CI]	No ACS (FP) [95% CI]	ACS (FN) [95% CI]	No ACS (TN) [95% CI]	
International data* ( <i>n</i> = 1 054 729 [1 047 881–1 061 578])	3.4% [2.8–4.0%]	10.8% [10.2–11.5%]	1.5% [0.8–2.2%]	84.3% [83.5–85.0%]	Beta
NL and B data** ( <i>n</i> = 862 960)	6.1% [4.0–7.9%]	8.1% [6.0–11.1%]	0.3% [0.0–2.5%]	85.5% [80.6–89.1%]	Beta
Combined data*** ( <i>n</i> = 1 054 729 [1 047 881–1 068 427])	6.8% [5.5–7.0%]	21.9% [21.6–23.3%]	0.3% [0.0–1.7%]	71.0% [69.6–71.7%]	Beta

Number of chest pain patients in the Netherlands and probabilities of ‘true positives’ (ACS positive referrals), ‘false positives’ (ACS negative referrals), ‘false negatives’ (ACS negative non-referrals) and ‘true negatives’ (ACS negative non-referrals), as well as the accompanying 95% CI and the distribution applied. Numbers and distributions are presented for three different scenarios: based on pooled analysis of international literature (international data), based on a Dutch/Belgian cohort study (NL and B data) and based on a combination of sources (combined data). As the international data are based on the largest set of patients, those were used in the base case analysis of this article.

\*References: Rosser *et al.* (1990), Klinkman *et al.* (1994), Svavarsdottir *et al.* (1996), Carroll *et al.* (2003), Nilsson *et al.* (2003), Bakx *et al.* (2005), Ruigomez *et al.* (2006), Koek *et al.* (2007), Verdon *et al.* (2008), Bosner *et al.* (2009), Yeh *et al.* (2010), Haasenritter *et al.* (2012), Soler *et al.* (2012b), Bruins Slot *et al.* (2013), Andersson *et al.* (2015), Frese *et al.* (2016).

\*\*References: Hoorweg *et al.* (2017). \*\*\*References: Klinkman *et al.* (1994), Svavarsdottir *et al.* (1996), Carroll *et al.* (2003), Nilsson *et al.* (2003), Bakx *et al.* (2005), Ruigomez *et al.* (2006), Koek *et al.* (2007), Verdon *et al.* (2008), Bosner *et al.* (2009), Yeh *et al.* (2010), Haasenritter *et al.* (2012), Soler *et al.* (2012b), Bruins Slot *et al.* (2013), Andersson *et al.* (2015), Frese *et al.* (2016), Hoorweg *et al.* (2017).

95% CI = 95% confidence interval; ACS = acute coronary syndrome; B = Belgium; FN = false negatives; FP = false positives; NL = the Netherlands; TN = true negatives; TP = true positives.



considered relevant after screening the title and abstract. Of the remaining 17 articles, three additional articles were eliminated after full text evaluation, since no relevant data were obtainable from these papers (Sox *et al.*, 1990; Soler *et al.*, 2012a; Haasenritter *et al.*, 2015). Two articles reported as references in the remaining 14 articles were added, resulting in a final number of 16 articles that were included (Supplementary Material Figure S1) (Rosser *et al.*, 1990; Klinkman *et al.*, 1994; Svavarsdottir *et al.*, 1996; Nilsson *et al.*, 2003; Carroll *et al.*, 2003; Bakx *et al.*, 2005; Ruigomez *et al.*, 2006; Koek *et al.*, 2007; Verdon *et al.*, 2008; Bosner *et al.*, 2009; Yeh *et al.*, 2010; Haasenritter *et al.*, 2012; Soler *et al.*, 2012b; Bruins Slot *et al.*, 2013; Andersson *et al.*, 2015; Frese *et al.*, 2016). In addition, data from regional and national Dutch and Belgian databases were used (Okkes *et al.*, 2005; Academic Hospital Maastricht, 2013; Catholic University Leuven, 2013; Dutch Heart Foundation, 2013; CBS, 2014a; 2014b; 2014c; 2014d). In Supplementary Material Table S1, all relevant data representing parts of the patient flow of patients with chest complaints, found in the selected articles and databases, are presented. All data were converted to absolute patient numbers in the Netherlands (Supplementary Material Table 2). Numbers of TPs, FPs, TNs and FNs were calculated (see Table 2). Besides the estimated patient numbers based on the two previously defined data sets (ie, base case ‘international data’ and the ‘NL and B data’), a third scenario (referred to as ‘combined data’) was defined. This data set was based on the international data, although one article was excluded because the health system and time setting in which this study was performed were considered not comparable with the current health system in the Netherlands (Rosser *et al.*, 1990). In addition, in this scenario the incidence of FNs was based on the ‘NL and B data’, as the incidence of FNs in the base case ‘international data’ seemed higher than observed in Dutch daily practice (Hoorweg *et al.*, 2017). All analyses were performed for each of these three scenarios, with the international data used as base case, as these data are based on the largest set of patients. The main results based on the ‘NL and B data’ and the ‘combined data’ are also presented in the text of this paper, the accompanying figures can be found in the supplementary web-only figures.

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### **Patient flow, GP’s sensitivity and specificity in current practice**

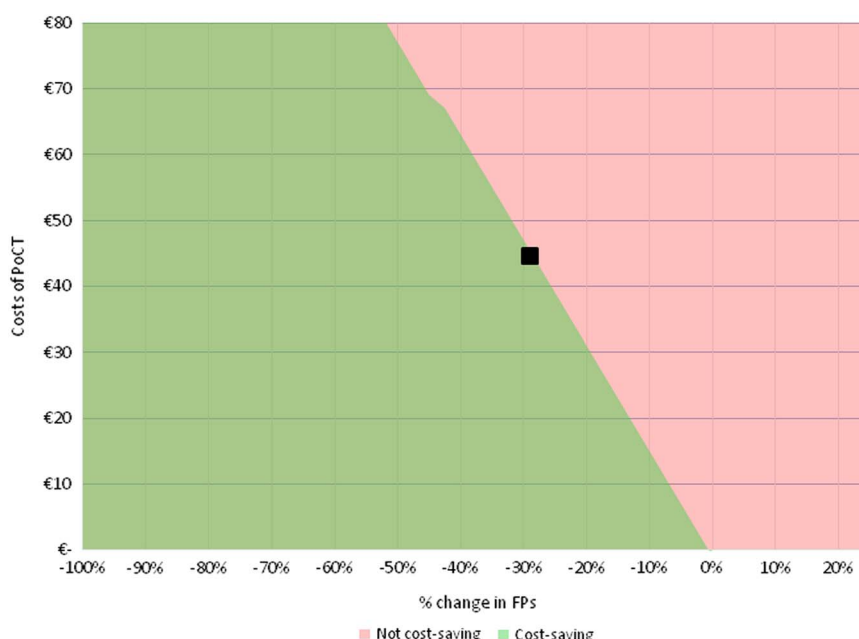
The results of the literature review indicate that annually in primary care in the Netherlands, 1 054 729 [95% CI 1 047 881–1 061 578] patients consult a GP for chest pain. The referral rate among these patients was found to be 14.2% [95% CI 14.0–14.4]. Sensitivity and specificity of a GPs judgement in the current setting (not aided by a CDR), are 69 and 89% for the ‘international’, 95 and 91% for ‘NL and B’, and 96 and 76% for the ‘combined data’, respectively.

### **Required reduction of ACS negative referrals (FPs) due to a future CDR, effect of further reduction of FPs**

When the cost price of a future PoCT is set at €45, the minimum required relative reductions in FPs for the PoCT strategy to become cost-saving are 29.0, 39.5 and 14.5% for the ‘international data’, the ‘NL and B data’, and the ‘combined data’, respectively (see Figure 1 and Supplementary Material Figures S2 and S3). In Table 3, the impact of a further relative reduction in FP rates on costs is shown for all three scenarios. For every additional absolute 10% reduction in %FPs, average additional cost savings per patient are €16 when using ‘international data’, €12 for ‘NL and B data’ and €31 for ‘combined data’ per chest pain patient.

### **Impact on health outcomes and costs**

The minimum required relative decrease in % FPs, as obtained from the two-way SA, was used as input into the PSA, while assuming that costs of the PoCT would remain unchanged (ie, €45). The results of this analysis on the average total costs (both per patient as well as the total costs in the Netherlands), are shown in Table 4. To visualize the constitution of those total costs, results are split up into costs that are attributable to TPs, FPs, FNs and TNs, and as the corresponding fraction of total costs. The result of the 10 000 Monte Carlo simulations is shown in Figure 2 for the ‘international data’, whereas results for the ‘NL and B data’ and ‘combined data’ are shown in Supplementary Material Figures S4 and S5, respectively.



**Figure 1** Two-way SA for ‘international data’. Deterministic two-way SA showing the combined effect of a relative reduction in ACS negative referrals (FPs, on x-axis), and of a variation in costs of a PoCT (on y-axis), on the difference in total costs between PoCT and current practice (ie, without PoCT). The analysis was performed based on the ‘international data’. When assuming that PoCT would only impact the %FPs and incur costs of the PoCT test (and leave all other model input parameters unaffected), a relative reduction of at least 29.0% in FPs is required to make the PoCT strategy become cost-saving (as represented by the black square, assuming a cost price of a PoCT test of €45.00). ACS=acute coronary syndrome; FPs=false positives; PoCT=point-of-care test; two-way SA=two-way sensitivity analysis.

**Table 3** Effect of a stepwise reduction of acute coronary syndrome (ACS) negative referrals (FPs) on costs

Change in %FPs	International data		NL and B data		Combined	
	Effect on costs	% effect	Effect on costs	% effect	Effect on costs	% effect
	Base case without POC: € 366.57		Base case without POC: € 485.01		Base case without POC: €722.46	
	Base case with POC: €411.71		Base case with POC: € 530.14		Base case with POC: €757.59	
-100%	€ - 110.20	- 30.1%	€ - 71.93	- 14.8%	€ - 286.05	- 39.6%
-90%	€ - 94.66	- 25.8%	€ - 60.22	- 12.4%	€ - 236.73	- 32.8%
-80%	€ - 79.13	- 21.6%	€ - 48.52	- 10.0%	€ - 205.41	- 28.4%
-70%	€ - 63.60	- 17.4%	€ - 36.81	- 7.6%	€ - 174.09	- 24.1%
-60%	€ - 48.07	- 13.1%	€ - 25.10	- 5.2%	€ - 142.77	- 19.8%
-50%	€ - 32.53	- 8.9%	€ - 13.40	- 2.8%	€ - 111.46	- 15.4%
-40%	€ - 17.00	- 4.6%	€ - 1.69	- 0.3%	€ - 80.14	- 11.1%
-30%	€ - 1.47	- 0.4%	€ 10.01	+ 2.1%	€ - 48.82	- 6.8%
-20%	€ 14.07	+ 3.8%	€ 21.72	+ 4.5%	€ - 17.50	- 2.4%
-10%	€ 29.60	+ 8.2%	€ 33.43	+ 6.9%	€ 13.81	+ 1.9%
0% (base case)	€ 45.13	+ 12.3%	€ 45.13	+ 9.3%	€ 45.13	+ 6.2%

This table shows the impact of steps of a 10% relative decrease in FP referrals (deterministic). At the top of the table, the costs for each of the base case scenarios is shown, depending on whether the PoCT is used. Absolute and relative effects are given for all three scenarios (‘international data’, ‘NL and B data’ and ‘combined data’ respectively). If no reduction in FPs is achieved (0% change = base case for all three scenarios where the PoCT is used) costs in all three scenarios will rise with the cost of a PoCT (€ 45.13), as compared with current daily practice where a PoCT is not used. ACS=acute coronary syndrome; B = Belgium; FP = false positives; NL = the Netherlands; PoCT = point-of-care test.

**Table 4** Costs per patient, converted to patient numbers in the Netherlands

Scenario	Referred				Not referred				Costs per patient	Total costs in the Netherlands
	ACS (TP)		No ACS (FP)		ACS (FN)		No ACS (TN)			
	Costs	%	Costs	%	Costs	%	Costs	%		
International data*	€194.00 (€116.35–293.90)	52.9%	€157.22 (€108.36–215.09)	42.9%	€0.28 (€0.12–0.50)	0.1%	€15.25 (€8.71–23.66)	4.2%	€366.71 (€268.07–483.72)	€385 736 500 (€282 739 424–510 189 897)
NL and B data**	€338.82 (€182.95–546.32)	71.3%	€120.88 (€72.67–184.89)	25.4%	€0.12 (€0.00–0.47)	0.0%	€15.40 (€8.76–23.87)	3.2%	€475.27 (€296.30–712.19)	€400 543 692 (€255 691 986–614 595 666)
Combined data***	€371.23 (€224.78–559.47)	52.5%	€322.99 (€221.85–445.24)	45.7%	€0.08 (€0.00–0.33)	0.0%	€12.79 (€7.31–20.00)	1.8%	€707.07 (€518.55–936.05)	€767 833 684 (€546 928 106–987 276 217)

This table presents where the average costs per patient are composed of, by splitting up those average costs into costs that are attributable to ACS positive referrals (TPs), ACS negative referrals (FPs), ACS positive non-referrals (FNs), and ACS negative non-referrals (TNs), as well as the accompanying percentage. Costs are given for all three base case scenarios ('international data', 'NL and B data' and 'combined data' respectively, without using PoCT), and based on the results of the probabilistic analysis.

\*References: Rosser *et al.* (1990), Klinkman *et al.* (1994), Svavarsdottir *et al.* (1996), Carroll *et al.* (2003), Nilsson *et al.* (2003), Bakx *et al.* (2005), Ruigomez *et al.* (2006), Koek *et al.* (2007), Verdon *et al.* (2008), Bosner *et al.* (2009), Yeh *et al.* (2010), Haasenritter *et al.* (2012), Soler *et al.* (2012b), Bruins Slot *et al.* (2013), Andersson *et al.* (2015), Frese *et al.* (2016).

\*\*Reference: Hoorweg *et al.* (2017). \*\*\*References: Klinkman *et al.* (1994), Svavarsdottir *et al.* (1996), Carroll *et al.* (2003), Nilsson *et al.* (2003), Bakx *et al.* (2005), Ruigomez *et al.* (2006), Koek *et al.* (2007), Verdon *et al.* (2008), Bosner *et al.* (2009), Yeh *et al.* (2010), Haasenritter *et al.* (2012), Soler *et al.* (2012b), Bruins Slot *et al.* (2013), Andersson *et al.* (2015), Frese *et al.* (2016), Hoorweg *et al.* (2017).

ACS = acute coronary syndrome; B = Belgium; FN = false negatives; FP = false positives; NL = the Netherlands; PoCT = point-of-care test; TN = true negatives; TP = true positives.

### Sensitivity of model outcome to changes in cost input parameters

The sensitivity of the model outcome to changes in individual cost input parameters, was measured using a one-way sensitivity analysis. The results are shown in tornado diagrams (see Figure 3 for 'international data', Supplementary Material Figures S6 and S7 for 'NL and B data' and 'combined data'). Results indicate that the model outcome (expressed as cost per patient) is robust to changes in input parameters in all three scenarios. In addition, in all three scenarios, the model outcome is most sensitive to changes in costs of the PoCT, while it is less sensitive to changes in costs of ambulance transportation to the hospital and costs of analysis at the coronary care unit among FPs. As each patient is assumed to undergo one consultation at the GP, and because the probability of correctly diagnosing ACS was assumed not to be affected by the PoCT, the model outcome was not affected by changes in those input parameters.

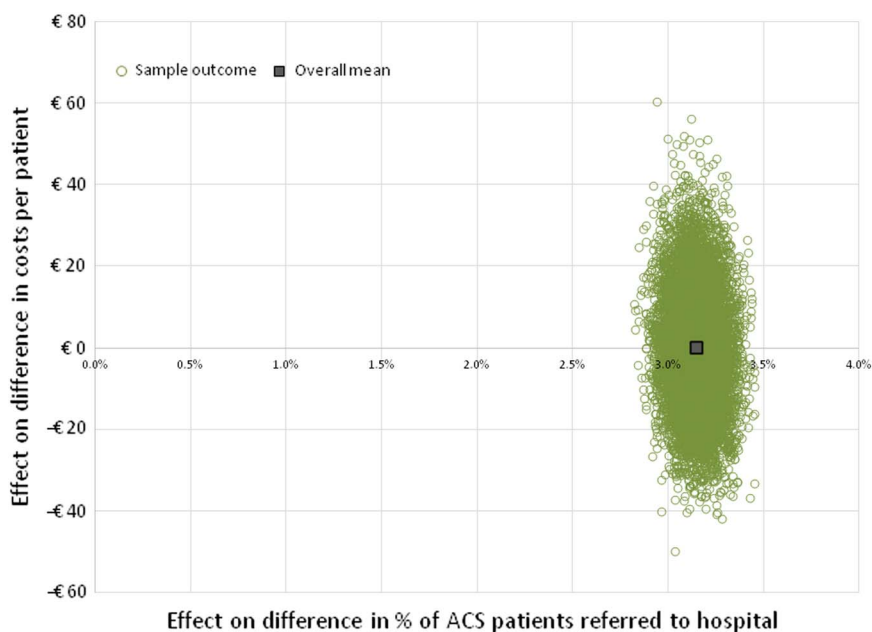
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### Discussion

#### Summary of main findings

We estimated that in the Netherlands (population 17 million) annually ~1 million patient contacts with GPs are about chest pain. In 14% of these contacts, direct referral to a cardiologist is made. Eventually, no more than a quarter of these referred patient is diagnosed with ACS (3.4% of all chest pain patients). As a result, 10.8% of all chest pain patients are referred while they eventually are diagnosed as not having ACS (FP). The estimated annual number of FPs in the Netherlands is 113 911, representing an economic burden of 162 million Euros. Improving triage of patients presenting with chest complaints to their GP could lead to a considerable reduction of FPs and thus to a reduction in both direct healthcare costs and patients' distress. We estimated the impact of a CDR incorporating a PoCT in this diagnostic process, on the accompanying costs and effects





**Figure 2** Incremental cost-effectiveness plane based on ‘international data’. This figure shows the result of 10000 model simulations (PSA), and the mean value, based on the international data. Costs of a PoCT are set at € 45, and reduction of ACS negative referrals (FPs) is assumed to be 29.0% (cost-neutral as compared with current practice, see Figure 1). ACS=acute coronary syndrome; FPs=false positives; PoCT=point-of-care test; PSA=probabilistic sensitivity analysis.

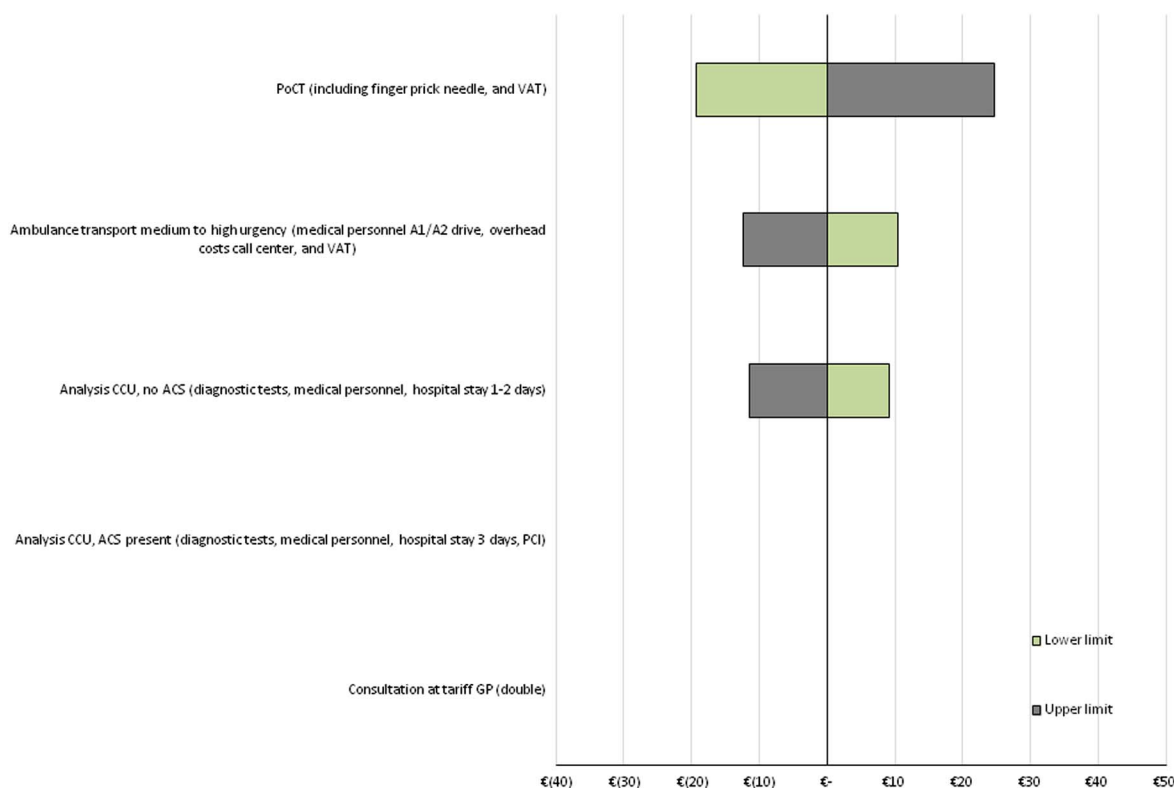
(ie, number of patients in whom the correct referral decision is made). When using an estimated cost price of a PoCT of €45, introduction of such test would be cost-saving if a relative reduction in FPs of at least 29% is achieved. This would imply a reduction in percentage of unnecessary referred patients from 10.8 to 7.7%. Such decrease would account for a cost saving of €47,106,755. Besides, for every 10% further reduction in %FPs, beyond the reduction of 29% where cost neutrality was reached, €16 per chest pain patient – referred or not referred – is saved, accounting for an annual saved amount of ~16 million Euros in the Netherlands. In the two alternative scenarios, a required reduction in %FPs of 39.5% for the ‘NL and B data’ and 14.4% for the ‘combined data’ was found. Such a reduction of FPs seems achievable when compared with results of similar studies in the field of suspected pulmonary embolism (Geersing *et al.*, 2012). Yet, a lower cost price of a PoCT can attenuate the required reduction in FPs. Halving a PoCT’s cost price to €22.50 leads to a minimum required reduction of FPs of only 14.5%

for the PoCT to become cost-saving, when based on the ‘international data’. Although the effect of preventing ACS negative referrals on societal costs has not been included in the current analysis, including those costs for both patients (and family or caregivers) would likely have increased the estimated cost savings that can be achieved by implementation of a PoCT-aided CDR.

### Sensitivity analysis

In the one-way model-sensitivity analysis, the model outcome proved to be robust for varying the model input parameters with –25% and +25% from the base case value. However, the starting point for variation in costs was based on a cost price for the PoCT of €45, which was based on the cost prognosis of a PoCT H-FABP test in development, and results may have been different when differently priced or different cardiac marker based PoCTs (eg, PoCT troponin) would have been used. Therefore, a wider range of costs was applied in the two-way sensitivity

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**Figure 3** Tornado diagram of one-way SA's for 'international data'. Tornado diagram showing the impact of changes in input parameters on the difference in costs, based on 'international data'. Costs of a PoCT are set at € 45, and the reduction of ACS negative referrals (FPs) is assumed to be 29.0% (cost-neutral situation as compared with current practice, see Figure 1). All input parameters were varied with 25% below and above the mean value. ACS = acute coronary syndrome; CCU = coronary care unit; FPs = false positives; PCI = percutaneous coronary intervention; one-way SA = one-way sensitivity analysis; PoCT = point-of-care test; VAT = value-added tax.

analysis, as this allows to apply the model to a wider range of PoCT cardiac markers for use in primary care.

### Strength and weakness

A strength of this study is that, to our knowledge, this study is the first to describe the possible financial benefit of a CDR incorporating a PoCT in chest pain patients presenting in primary care. We synthesized all available and relevant evidence, from different studies and countries, to make the best possible estimation of prevalence of chest complaints and referral rates in primary care. Moreover, we repeated the analysis using three different data sets for patient numbers of chest

pain patients in primary care in the health economic model. As the base case model outcome is based on several international studies, it is likely that the results regarding the effectiveness can be generalized to other countries. However, as costs may differ strongly between countries, country-specific cost estimates are required to make reliable per-country calculations.

In our analyses it was assumed that the test is performed in all patients. However, when a CDR is already positive after only scoring a patient's clinical findings by the GP, the patient will most likely be immediately referred without performing a PoCT. Therefore, as the use of PoCT will likely be more sophisticated in daily clinical practice, the costs of PoCT have (conservatively) been

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overestimated in the current analysis. On the other hand, wide availability of a validated PoCT in the future could lower the threshold for using this test in daily practice, as expected based on previous research (Kip *et al.*, 2017).

Still, some uncertainties remain. Despite the fact that patient numbers and referral rates are based on a thorough review of available literature, the number of studies that could eventually be included was limited for some model input parameters. In addition, some of those studies were relatively small. Although it might have been most straightforward to use a Dirichlet distribution to estimate all four probabilities (TP, FP, TN and FN) simultaneously, this would have required to use a small number of patients across this distribution, thereby largely overestimating the uncertainty in model outcomes. Therefore, a Beta distribution was used instead, which allows a two-step approach. First, the probability of a patient being referred was estimated (which was based on a large number of patients), followed by estimating the probability of being either TP or FP (among referrals), or TN or FN (among non-referrals). As we assumed that the number of patients with and without ACS had to remain constant, patients could only switch from TP to FN, and from TN to FP (and vice versa). Consequently, we could not simultaneously incorporate uncertainty in the number of patients that either have or do not have ACS. However, as the result of Monte Carlo simulations shows that the uncertainty in costs is expected to be limited, and because the analysis has been performed for three different scenarios, we consider it unlikely that incorporating this uncertainty would have changed the conclusions.

Furthermore, a validated CDR does not yet exist and until such a diagnostic tool is available, the actual diagnostic performance of a CDR incorporating a POCT remains uncertain. In addition, an issue that could not be included in the current model (as it was considered too complex for the scope of this analysis), is the possible presence of other severe diseases causing chest pain in patients where ACS is excluded. However, in recent studies, nearly all FPs are diagnosed with not life threatening diseases (Bruins Slot *et al.*, 2013; Hoorweg *et al.*, 2017). Another possible limitation may be caused by the fact that the number of FNs was assumed to remain constant in the current analysis. In current practice, patients

that are initially not referred but are eventually diagnosed as having ACS (in the following days after assessment by a GP) might be under registered. Moreover, a PoCT could reduce FN referrals, which likely leads to health gain. In addition, those FN rates were found to be relatively high in our ‘international data’ set, but appear to be much lower in recent studies (Hoorweg *et al.*, 2017), and based on the authors’ experience in daily clinical practice. Therefore, the ‘combined’ data set is assumed to provide an estimate better reflecting current practice in the Netherlands, and in many other countries with similar healthcare systems. However, because of the assumptions that had to be made, the ‘international’ data were conservatively chosen as base case scenario in the current analysis.

## Conclusion

If the use of a CDR including a PoCT can reduce unnecessary referrals of chest pain patients to secondary care, this could considerably reduce healthcare costs. Our study provides insights in the minimum requirements (regarding the relative reduction in FPs, as well as the costs of the PoCT) for the CDR strategy including a PoCT to become cost-saving.

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## Conflicts of Interest

J.G. is chief scientific officer (CSO) at FABPulous BV. The remaining authors report no conflicts of

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interest. The authors alone are responsible for the content and writing of the paper.

## Supplementary material

To view supplementary material for this article, please visit <https://doi.org/10.1017/S146342361700069X>

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