

# Cognitive impairment in depression: a systematic review and meta-analysis

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# Cognitive impairment in depression: a systematic review and meta-analysis

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**Background.** This review aimed to address the question of whether cognitive impairment should be considered a core feature of depression that may be a valuable target for treatment.

**Method.** We conducted a systematic review and meta-analysis of cognitive function, assessed with a single neuropsychological test battery, the Cambridge Neuropsychological Test Automated Battery (CANTAB), in patients with depression during symptomatic and remitted states. Inclusion of studies comparing patients remitted from depression and controls enabled us to investigate whether cognitive impairment persists beyond episodes of low mood in depression.

**Results.** Our meta-analysis revealed significant moderate cognitive deficits in executive function, memory and attention in patients with depression relative to controls (Cohen's *d* effect sizes ranging from  $-0.34$  to  $-0.65$ ). Significant moderate deficits in executive function and attention (Cohen's *d* ranging from  $-0.52$  to  $-0.61$ ) and non-significant small/moderate deficits in memory (Cohen's *d* ranging from  $-0.22$  to  $-0.54$ ) were found to persist in patients whose depressive symptoms had remitted, indicating that cognitive impairment occurs separately from episodes of low mood in depression.

**Conclusions.** Both low mood and cognitive impairment are associated with poor psychosocial functioning. Therefore, we argue that remediation of cognitive impairment and alleviation of depressive symptoms each play an important role in improving outcome for patients with depression. In conclusion, this systematic review and meta-analysis demonstrates that cognitive impairment represents a core feature of depression that cannot be considered an epiphenomenon that is entirely secondary to symptoms of low mood and that may be a valuable target for future interventions.

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**Key words:** Attention, CANTAB, cognition, depression, executive function, memory.

## Introduction

Cognitive impairment is frequently observed in patients suffering from depression and is associated with poor response to treatment (Potter *et al.* 2004; Story *et al.* 2008; Roiser *et al.* 2012). Impaired cognition has been estimated to occur in around two-thirds of depressed patients (Abas *et al.* 1990; Butters *et al.* 2004; Afridi *et al.* 2011). Impaired ability to think, concentrate or make decisions is a DSM-IV-TR (APA, 2000) diagnostic criterion for major depressive episode. Consistent with this, several systematic reviews have demonstrated cognitive deficits in patients suffering from depression (Burt *et al.* 1995; Veiel, 1997;

Zakzanis *et al.* 1998; Stefanopoulou *et al.* 2009; Snyder 2013), including first-episode patients (Lee *et al.* 2012).

Impairments in cognition have been found to persist beyond acute episodes of depression, and between one-third and one-half of remitted depressed patients are thought to be affected by cognitive deficits (Abas *et al.* 1990; Bhalla *et al.* 2006; Reppermund *et al.* 2009). Furthermore, one study revealed that 94% of patients who had cognitive impairment while depressed continued to experience deficits in cognition when remitted from depression (Bhalla *et al.* 2006).

To our knowledge, to date, only two groups have reviewed cognitive function in patients remitted from depression (Hasselbalch *et al.* 2011; Bora *et al.* 2013). The review by Hasselbalch *et al.* (2011) included 500 remitted patients (and 472 controls) and revealed impaired cognitive performance in nine of the 11 included studies. Their review also assessed the association between cognitive function and other clinical

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features such as residual depressive symptoms and current medication status. However, drawbacks of this review relate to the large number of different cognitive tests that were used across studies and the lack of implementation of standardized effect sizes to reflect magnitude of impairment. Meanwhile, the review by Bora *et al.* (2013) included 895 remitted patients (and 997 controls) from 27 studies and, using standardized effect sizes, revealed cognitive deficits in a composite measure of global cognition, in individual cognitive domain composites and in a subset of specific tasks. The review also separately assessed cognitive function in early-onset and late-onset patients and included a meta-regression to uncover the influence of other clinical and demographic factors on cognitive performance. Again, a minor drawback of this review is that task-specific analyses were limited to a subgroup of cognitive tests for which there were sufficient data; therefore, cognitive domain and global cognition meta-analyses necessarily included results from a variety of cognitive tests. A review of the longitudinal course of cognitive function in depression revealed that improvements in mood were most closely related to improvements in verbal memory, verbal fluency and psychomotor speed, whereas attention and executive function remained impaired across treatment (Douglas & Porter, 2009).

Our aim was to conduct a systematic review and meta-analysis to investigate the degree of cognitive impairment in patients with depression during symptomatic and remitted states, focusing on studies that used a single neuropsychological test battery, the Cambridge Neuropsychological Test Automated Battery (CANTAB). Our rationale for including only CANTAB studies was to enable assessment of a broad range of cognitive domains but with consistent tasks implemented across reviewed studies, thereby ensuring interstudy homogeneity. We predicted that cognitive deficits would be observable in both depressed and remitted states.

## Method

### Systematic review

Studies were identified by searching PubMed and Google Scholar using the following search terms: 'Cambridge neuropsychological test automated battery' or 'CANTAB' and any CANTAB test name (e.g. 'Spatial Span') or its acronym ('SSP') and 'depression' or 'depressed' during the period from 1980 to December 2012. The CANTAB neuropsychological tests included in the search involved the domains of executive function, memory, attention and reaction time, as follows.

### Executive function

*(One Touch) Stockings of Cambridge (OTS/SOC; Owen et al. 1990)*. This task was derived from the Tower of London test and assesses visual planning, reasoning and impulsivity. Outcome measures analysed were the number/percentage correct or number of moves above the minimum [for all problems or difficult (four/five-move) problems].

*Spatial Working Memory (SWM; Owen et al. 1995)*. This self-ordered search task is based on foraging behaviour and assesses working memory and strategy use. Participants search for tokens without returning to previous token locations. Outcome measure analysed was between-search errors.

*Intra-Extra Dimensional Set Shift (IED; Rogers et al. 1999)*. This test of cognitive flexibility, analogous to the Wisconsin Card Sorting Test (WCST), has multiple stages segregating cognitive processes that assess rule learning, rule reversal and attentional set-shifting. Outcome measures analysed were total errors, extra-dimensional shift errors (adjusted) or stages completed.

*Spatial Span (SSP; Kempton et al. 1999)*. This is a task of spatial short-term memory based on the Corsi block-tapping task. Outcome measure analysed was spatial span.

### Memory

*Delayed Matching to Sample (DMS; Robbins et al. 1994)*. In this test participants remember the visual features of a complex, abstract target stimulus and select it from a choice of four target patterns after a variable delay. Outcome measures analysed were total/percentage correct (for all trials or 12-s delay trials).

*Paired Associates Learning (PAL; Sahakian et al. 1988)*. In this test participants learn the locations of a progressively increasing number of abstract stimuli. Outcome measures analysed were total errors (adjusted) or first trials correct.

*Pattern Recognition Memory (PRM; Owen et al. 1995)*. This is a two-forced-choice test of abstract visual pattern recognition memory. Outcome measures analysed were total/percentage correct.

*Spatial Recognition Memory (SRM; Owen et al. 1995)*. This two-forced-choice discrimination paradigm tests spatial recognition memory. Outcome measures analysed were total/percentage correct.

### Attention

*Rapid Visual Information Processing (RVP; Sahakian et al. 1989)*. This is a continuous performance test that assesses sustained attention, signal detection and impulsivity. Participants monitor a stream of single digits for three-digit target sequences. Outcome measures analysed were target sensitivity or total hits/omissions.

### Reaction time

*Reaction Time (RTI; Sahakian et al. 1993)*. This is a test of simple and five-choice reaction time. Outcome measure analysed was five-choice reaction time.

### Inclusion criteria

The inclusion criteria for studies were: (1) used DSM or ICD criteria to diagnose major depressive disorder; (2) included a healthy control group; (3) used CANTAB to assess cognitive function in currently depressed patients and/or remitted depressed patients; and (4) reported sufficient data to estimate Cohen's  $d$  effect sizes, that is the group mean and either standard deviation or standard error data (and number of subjects in each group) were available for both patients and controls.

Our search revealed 24 studies including 784 currently depressed patients (and 727 controls) and six studies including 168 remitted depressed patients (and 178 controls) that met our inclusion criteria (see Table 1). The criteria for remitted depression varied across studies and are shown in Table 1.

### Meta-analysis

Meta-analysis was performed using Review Manager (RevMan, 2011). For each study, Cohen's  $d$  effect sizes (Cohen, 1988) were calculated as the mean difference between test performance scores for patients compared to controls divided by the pooled standard deviation; negative effect sizes reflected deficits compared to controls. Subsequently, for each test, effect sizes were weighted using the inverse variance method within a random-effects model and pooled across all studies with available data. Pooled effect sizes were reported for tests only when data from three or more studies were available. In addition to meta-analyses for currently depressed patients *versus* controls and remitted depressed patients *versus* controls, a separate subanalysis was conducted for currently depressed patients who were unmedicated at the time of assessment *versus* controls. There were insufficient studies of unmedicated remitted depressed patients to include a subanalysis of this population. Influenced by Cohen's convention regarding the magnitude of effect sizes (Cohen, 1988), a Cohen's  $d$  effect size in

the range 0.2–0.35 was considered small, in the range 0.35–0.65 moderate and >0.65 large. Statistical inferences were made based upon analysis of 95% confidence intervals (CIs).

## Results

### *Profile of cognitive deficits in currently depressed patients*

Cohen's  $d$  effect sizes were calculated based on data from 24 studies that used CANTAB tests in 784 currently depressed patients and 727 controls. Fig. 1 shows the weighted, pooled Cohen's  $d$  effect sizes for the comparison between depressed patients and healthy controls (black bars), and Table 2 presents detailed meta-analysis results.

Currently depressed patients showed significant moderate deficits compared to healthy controls across the cognitive domains of executive function (Cohen's  $d$  ranged from  $-0.34$  to  $-0.54$ ), memory (Cohen's  $d$  ranged from  $-0.41$  to  $-0.50$ ) and attention (Cohen's  $d$  was  $-0.65$ ), and there was no significant deficit in reaction time (Cohen's  $d$  was  $-0.07$ ). The non-significant finding for reaction time should be treated with caution because the results seem to have been affected by one study for which depressed patients showed significantly superior performance to controls. Indeed, when this study was excluded, currently depressed patients showed a nearly significant small deficit in reaction time compared to controls ( $d = -0.32$ , 95% CIs  $-0.59$  to  $-0.05$ ). Supplementary Fig. S1 (available online) presents forest plots depicting performance of currently depressed patients relative to controls.

### *Subanalysis: profile of cognitive deficits in unmedicated currently depressed patients*

Cohen's  $d$  effect sizes were calculated based on data from eight studies that used CANTAB tests in 271 currently depressed patients who were unmedicated at the time of assessment and 267 controls. There were sufficient data to calculate weighted, pooled effect sizes for all executive function tasks, all memory tasks, and for the task of attention; insufficient data were available to calculate a weighted, pooled effect size for the reaction time task. Table 2 presents detailed meta-analysis results.

Unmedicated currently depressed patients showed significant moderate deficits compared to healthy controls on one executive function task (SWM; Cohen's  $d$  was  $-0.46$ ), two memory tasks (DMS and PRM; Cohen's  $d$  ranged from  $-0.33$  to  $-0.36$ ) and the attention task (RVP; Cohen's  $d$  was  $-0.59$ ). Although negative Cohen's  $d$  effect sizes (ranging from  $-0.06$  to  $-0.49$ ) were recorded for all remaining tasks, the

**Table 1.** Study characteristics and patient demographics for currently depressed and remitted depressed comparisons

First author	Year	Currently depressed patients					Controls		Notes
		<i>n</i> (female)	Age (years)	Diagnostic criteria	Depression symptoms	Medication status	<i>n</i> (female)	Age (years)	
Beats	1996	24 (12)	72.0±5.9	DSM-III-R	HAMD-x 29.6±5.1; MADRS 40.3±7.2	Twenty-one medicated, three medication free	15 (9)	69.3±6.6	Minimum age of 60
Boeker	2012	28 (13)	39.7±11.4	DSM-x; HAMD-21 ≥24; BDI ≥24	HAMD-21 28.5±7.0; BDI 25.9±8.2	Nineteen medicated, nine medication free	28 (13)	35.0±7.4	
Braw	2011	25 (14)	54.0±0.9	DSM-IV; HAMD-17>14	HAMD-17 31.3±1.3; BDI 30.8±1.4	All unmedicated for 1 month prior to testing	25 (17)	54.2±0.9	Late adulthood group aged 46–65
		30 (16)	35.0±1.0	DSM-IV; HAMD-17>14	HAMD-17 32.5±1.1; BDI 33.5±1.4	All unmedicated for 1 month prior to testing	30 (16)	34.5±1.1	Middle adulthood group aged 25–45
		30 (20)	17.1±0.5	DSM-IV; CDRS ≥40	CDRS-R 67.5±2.0; BDI 32.6±1.3	All unmedicated for 1 month prior to testing	30 (18)	17.5±0.6	Young adult group aged <25
Cannon	2009	18 (11)	31±11	DSM-IV	MADRS 22±5.3; IDS-C 27±6.5	All unmedicated (of whom 11 treatment naïve)	19 (11)	31±8.5	Aged 18–55
Elliott	1996	28 (19)	49.9±1.7	DSM-III-R	HAMD-x 22.4±0.8; MADRS 34.0±1.1	All medicated	22 (15)	48.1±1.2	Aged 40–70
Elliott	1997	6 (1)	34.7 (21–48)	DSM-IV	HAMD-x 23.8 (20–29); MADRS 35.3 (x-39)	Five medicated, one unmedicated	6 (1)	31.0 (18–55)	
Erickson	2005	20 (10)	37.2±11.9	DSM-IV	MADRS 25.4±7.1	All unmedicated for 3 weeks prior to testing (of whom four medication naïve)	Matched (not stated)	Matched (not stated)	All had illness onset before age 40
Grant	2001	48	39.0±10.4	DSM-IV	HAMD-17 16.7±5.4	All unmedicated patients for 28 days prior to testing	31	40.2±9.7	Demographics are for a larger sample from which these subjects were drawn
Heinzel	2010	20 (11)	40.0±9.9	DSM-IV; HAMD-21 ≥24	HAMD-21 33.1±7.1; BDI 29.9±4.9	All unmedicated for 1 week prior to testing	29 (21)	35.3±7.3	
Kyte	2005	30 (18)	15.3±2.5	K-SADS-PL	HAMD-x 10.9±6.8	Medicated and unmedicated adolescents	49 (29)	15.2±2.1	
Lyche	2010	37 (23)	44.2±12.3	DSM-IV	BDI 21.4±11.1	Thirteen medicated, 24 unmedicated	91 (63)	35.8±12.0	
Maalouf	2010	20 (16)	34.2±9.4	DSM-IV; HAMD-25 ≥17	HAMD-5 24.8±5.8	All medicated	28 (19)	31.9±9.4	
Maalouf	2011	20 (17)	15.3±1.6	DSM-IV; K-SADS-PL	CDRS 58.6±10.9	Thirteen medicated, seven unmedicated	17 (9)	15.2±1.8	
Matthews	2008	14 (14)	14.5±1.2	ICD-10; CAPA-C	MFQ 41.3±10.4	All medication naïve	14 (14)	14.4±1.0	
Michopoulos	2008	40 (40)	52.7±10.8	DSM-IV-TR	HAMD-17 20.0±4.0	All medicated	20 (20)	49.8±12.7	

Michopoulos	2006	11 (11) 11 (11)	50.9±10.5 47.8±12.3	DSM-IV DSM-IV	HAMD-x 20.8±3.1 HAMD-x 18.7±4.2	All medicated All medicated	11 (11)	52.8±14.1	Melancholic subgroup Non-melancholic subgroup	
Murphy	2003	27 (14)	38.9±9.7	DSM-IV	HAMD-x 23.6±4.2; MADRS 34.3±5.4	Twenty-six medicated, one unmedicated	23 (12)	39.1±10.8		
O'Brien	2004	61 (48)	73.9±6.7	DSM-IV; MADRS ≥20	MADRS 30.7±7.1	Mostly medicated (numbers not stated)	40 (30)	73.3±6.7	Aged over 60	
Porter	2003	44 (29)	32.9±10.6	DSM-IV	HAMD-17 21.1±4.4; MADRS 28.9±5.5; BDI 27.9±10.2	All unmedicated (of whom 26 medication naïve) for 6 weeks prior to testing	44 (29)	32.3±11.4		
Purcell	1997	20 (12)	37.5 (18–52)	DSM-IV	HAMD-24 22.6±5.6	Twelve medicated, eight unmedicated for 2 months prior to testing	20 (12)	37.2 (21–60)		
Reppermund	2009	53 (28)	43.5±8.0	DSM-IV	HAMD-x 25.1±5.1	Fifty medicated, three unmedicated	13 (7)	46.4±9.5		
Swainson	2001	37	60.8±8.6	DSM-IV	HAMD-x 21.4±6.2	Not stated	39	64.4±8.5		
Sweeney	2000	58 (39)	32.3±9.1	DSM-IV	HAMD-17 21.6±4.3	Medicated patients	51 (39)	36.3±9.7		
Taylor	2007	22 (17)	38.6±8.1	DSM-IV	MADRS 25.5±7.5	Unmedicated patients	25 (18)	34.8±8.8		
Tavares										
Tsaltas	2010	15 (15) 15 (15)	47.8±11.7 48.5±11.2	DSM-IV-TR DSM-IV-TR	HAMD-24 27.6±5.6 HAMD-24 31.9±6.5	All medicated All medicated	15 (15)	49.3±11.6	Non-referred subgroup Referred subgroup	
Remitted patients										
							Controls			
First author	Year	<i>n</i> (female)	Age (years)	Diagnostic criteria	Euthymia definition	Depression symptoms	Medication status	<i>n</i> (female)	Age (years)	Notes
Beats	1996	19 (10)	73.6±5.4	DSM-III-R	MADRS <10	HAM-D 4.7±2.6; MADRS 6.5±4.5	Mostly medicated	15 (9)	69.3±6.6	
Clark Clark	2005 <sup>a</sup> 2005 <sup>b</sup>	15 (11)	45.2±10.9	DSM-IV	HAMD-x <9	HAMD-x 2.1±2.9	Six medicated, nine unmedicated	46 (23)	39.2±12.2	
Herrera- Guzman	2010	60	20–50	DSM-IV	HAMD-17 <6	HAMD-17 0.7±0.2	All unmedicated	37	20–50	
Maalouf	2011	20 (15)	15.4±1.3	DSM-IV; K-SADS-PL	CDRS ≤28	CDRS 23.7 ±10.9	Thirteen medicated, seven unmedicated	17 (9)	15.2±1.8	

Table 1 (cont.)

First author	Remitted patients				Controls					
	Year	n (female)	Age (years)	Diagnostic criteria	Euthymia definition	Depression symptoms	Medication status	n (female)	Age (years)	Notes
O'Brien	2004	26	Not stated	DSM-IV	MADRS <8	Not stated	Not stated	40	78.3±6.7	Aged over 60
Weiland-Fiedler	2004	28 (18)	37.8±12.2	DSM-IV	MADRS <6	MADRS 2.1±2.3	All unmedicated for 3 months prior to testing	23 (11)	35.7±10.4	

DSM-x, *Diagnostic and Statistical Manual of Mental Disorders* (APA, 2000); HAMD-17/21/24/25/-x, Hamilton Depression Rating Scale (17-/21-/24-/25-item/unstated version) (Hamilton, 1960); MADRS, Montgomery-Asberg Depression Rating Scale (Montgomery & Asberg, 1979); BDI, Beck Depression Inventory (Beck, 1961); K-SADS-PL, Schedule for Affective Disorders and Schizophrenia for School-Age Children – Present and Lifetime Version (Kaufman et al. 1997); CDRS, Children's Depression Rating Scale (Foznanski et al. 1979); CAPA-C, Child and Adolescent Psychiatric Assessment – Child Version (Angold & Costello, 1995); MFQ, Mood and Feelings Questionnaire (Angold et al. 2002).

Values given as mean±standard deviation or mean (range).

95% CIs crossed zero in all cases. Supplementary Fig. S2 presents forest plots depicting performance of unmedicated currently depressed patients relative to controls.

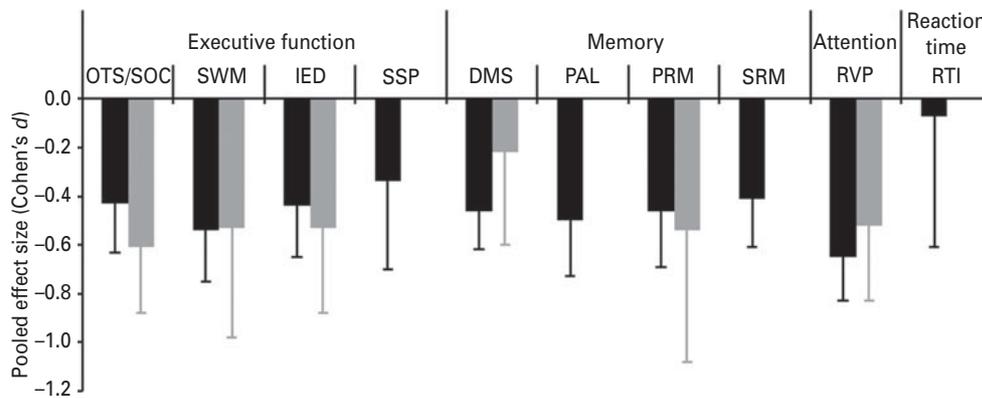
### Profile of cognitive deficits in remitted depressed patients

Cohen's *d* effect sizes were based on data from six studies that used CANTAB tests in 168 remitted depressed patients and 178 controls. There were sufficient data to calculate weighted, pooled effect sizes for three (out of four) tasks in the domain of executive function, two (out of four) tasks in the domain of memory, and for the task of attention; insufficient data were available to calculate a weighted, pooled effect size for the reaction time task. Fig. 1 shows the weighted, pooled Cohen's *d* effect sizes for the comparison between depressed patients and healthy controls (grey bars), and Table 2 presents detailed meta-analysis results.

Patients remitted from depression showed significant moderate deficits compared to healthy controls across the cognitive domains of executive function (Cohen's *d* ranged from  $-0.53$  to  $-0.61$ ) and attention (Cohen's *d* was  $-0.52$ ). There was a tendency towards moderate deficits in the domain of memory (Cohen's *d* ranged from  $-0.22$  to  $-0.54$ ). Although the 95% CIs crossed zero in both cases, they only just crossed zero for PRM (95% CIs were from  $-1.08$  to  $0.01$ ). Supplementary Fig. S3 presents forest plots depicting performance of currently depressed patients relative to controls.

### Discussion

Our systematic review and meta-analysis revealed that impairments in cognitive function, assessed with a single neuropsychological test battery (CANTAB), were exhibited by currently depressed patients and by patients remitted from depression. Current depression was associated with significant moderate deficits across all tasks within the domains of executive function, memory and attention, with the exception of the SSP task of executive function, for which there was a tendency towards a moderate deficit. Although the systematic review and meta-analysis revealed no reaction time deficit in currently depressed patients, exploratory reanalysis excluding one anomalous study (in which depressed patients showed significantly superior performance relative to controls) revealed a tendency towards a small deficit in reaction time. Analysis of only unmedicated currently depressed patients showed a significant moderate deficit in the domain of attention and significant small and moderate deficits in some, but not all, tasks within the



**Fig. 1.** Pooled, weighted Cohen's *d* effect sizes reflecting the performance of currently depressed patients (black bars) and remitted depressed patients (grey bars) compared to healthy controls on tasks of executive function [OTS/SOC, (One Touch) Stockings of Cambridge; SWM, Spatial Working Memory; IED, Intra-Extra Dimensional Set Shift; SSP, Spatial Span], memory (DMS, Delayed Matching to Sample; PAL, Paired Associates Learning; PRM, Pattern Recognition Memory; SRM, Spatial Recognition Memory), attention (RVP, Rapid Visual Information Processing) and reaction time (RTI, Reaction Time). Error bars represent 95% confidence intervals (CIs).

domains of executive function and memory. Meanwhile, remitted depressed patients showed significant moderate deficits within the domains of executive function and attention. However, in the domain of memory, remitted depressed patients showed only a tendency towards small/moderate deficits. In summary, our systematic review and meta-analysis demonstrated that cognitive impairment, particularly affecting the domains of executive function and attention, is a core feature of depression that persists during remission in the absence of clinically relevant symptoms of low mood.

The present systematic review and meta-analysis included only studies that had used CANTAB tasks to assess cognitive function in symptomatic or remitted depressed patients relative to controls. To our knowledge, this is the first systematic review and meta-analysis that has focused on studies using a single neuropsychological test battery. The magnitudes of cognitive deficits recorded in the current investigation are broadly in line with those that have been recorded previously. However, our finding of a non-significant deficit in reaction time in currently depressed patients relative to controls contrasted notably with the literature. Nevertheless, following exclusion of one anomalous result, a tendency towards a small deficit on the RTI task was recorded, and the size of this deficit (Cohen's  $d=0.32$ ) was similar to the deficit recorded on the psychomotor speed composite (Cohen's  $d=0.33$ ) in the Snyder (2012) meta-analysis.

Impaired cognitive functioning has been linked with poor response to antidepressant treatment (Potter *et al.* 2004; Story *et al.* 2008). However, the potential clinical relevance of cognitive deficits in depression also depends upon their impact on psychosocial functioning.

Impaired psychosocial functioning is a core feature of depression (Weissman *et al.* 2010). It persists in up to 60% of individuals with depression even after mood symptoms of depression have remitted (Jaeger *et al.* 2006), indicating that severity of depressive symptoms cannot fully account for impaired functional ability. For example, patients with subsyndromal depressive symptoms have been found to manifest similar levels of psychosocial dysfunction to those of patients with clinically relevant symptoms (Judd *et al.* 1996). One possible explanation is that persisting cognitive impairments may contribute to poor quality of life and psychosocial functioning in patients whose depressive symptoms have remitted. In support of this, psychosocial functioning has been shown to be associated with performance on measures of attention, executive function, paired associates learning and visuospatial ability in depression (Jaeger *et al.* 2006). Importantly, the association between cognitive deficits and poor psychosocial functioning has been shown to remain significant even when taking into account residual, subclinical depressive symptoms (Jaeger *et al.* 2006).

Another study revealed that severity of cognitive impairment and severity of low mood associate independently with different measures of psychosocial functioning (McCall & Dunn, 2003). Furthermore, in bipolar disorder, psychosocial functioning has been shown to be predicted by both cognition and residual depressive symptoms (Mur *et al.* 2009; Solé *et al.* 2012).

Overall, these findings suggest that remediation of cognitive impairment and alleviation of depressive symptoms may both be involved in improving psychosocial functioning in depression. We therefore argue that cognitive impairment in depression is clinically relevant and may be a valuable target for intervention.

**Table 2.** Meta-analysis results

Task	No. patients	No. controls	No. studies	<i>d</i>	95% CI	Z	<i>p</i>	Q	<i>I</i> <sup>2</sup> (%)
Currently depressed patients									
OTS/SOC	557	484	16	-0.43	-0.63 to -0.24	4.32	<0.0001	43.33	56
SWM	567	521	15	-0.54	-0.75 to -0.33	4.98	<0.00001	43.92	64
IED	578	566	16	-0.44	-0.65 to -0.23	4.07	<0.0001	52.97	64
SSP	273	217	8	-0.34	-0.70 to 0.01	1.92	0.06	24.19	71
DMS	423	342	12	-0.46	-0.62 to -0.29	5.37	<0.00001	13.52	19
PAL	321	279	9	-0.50	-0.73 to -0.26	4.17	<0.0001	18.43	46
PRM	402	347	12	-0.46	-0.69 to -0.23	3.89	0.0001	25.55	57
SRM	445	371	13	-0.41	-0.61 to -0.22	4.19	<0.0001	24.38	43
RVP	228	236	7	-0.65	-0.83 to -0.46	6.75	<0.00001	3.90	0
RTI	157	135	4	-0.07	-0.61 to 0.46	0.27	0.79	14.33	79
Unmedicated currently depressed patients									
OTS/SOC	191	174	4	-0.28	-0.68 to 0.11	1.40	0.16	17.21	71
SWM	231	218	6	-0.46	-0.84 to -0.09	2.43	0.02	25.85	73
IED	171	166	4	-0.09	-0.46 to 0.28	0.49	0.62	14.00	64
SSP	82	69	3	-0.06	-0.66 to 0.54	0.20	0.84	6.18	68
DMS	126	112	4	-0.36	-0.62 to -0.10	2.71	0.007	2.67	0
PAL	106	89	3	-0.49	-1.22 to 0.23	1.33	0.18	11.02	82
PRM	146	132	5	-0.33	-0.61 to -0.04	2.23	0.03	5.42	26
SRM	125	112	4	-0.29	-0.75 to 0.17	1.22	0.22	8.62	65
RVP	123	124	3	-0.59	-0.84 to -0.33	4.50	<0.00001	1.38	0
Remitted depressed patients									
OTS/SOC	125	109	4	-0.61	-0.88 to -0.34	4.47	<0.00001	2.80	0
SWM	114	100	3	-0.53	-0.98 to -0.07	2.28	0.02	5.00	60
IED	62	84	3	-0.53	-0.88 to -0.18	2.95	0.003	1.51	0
DMS	74	80	3	-0.22	-0.60 to 0.15	1.16	0.24	2.69	26
PRM	73	78	3	-0.54	-1.08 to 0.01	1.92	0.05	5.24	62
RVP	123	123	4	-0.52	-0.83 to -0.21	3.31	0.0009	3.87	22

*d*, Weighted, pooled Cohen's *d* effect size; CI, confidence interval; Q, heterogeneity; *I*<sup>2</sup>, percentage of total variability due to heterogeneity; OTS/SOC, (One Touch) Stockings of Cambridge; SWM, Spatial Working Memory; IED, Intra-Extra Dimensional Set Shift; SSP, Spatial Span; DMS, Delayed Matching to Sample; PAL, Paired Associates Learning; PRM, Pattern Recognition Memory; SRM, Spatial Recognition Memory; RVP, Rapid Visual Information Processing; RTI, Reaction Time.

Although there are relatively few published studies assessing the cognitive enhancing effects of pharmacological treatments in depression, one potential augmentation therapy is the wakefulness-promoting agent modafinil. Indeed, 4-week adjunctive treatment with modafinil was shown to improve performance on a task of executive function in currently depressed patients with only partial response to antidepressant therapy (DeBattista *et al.* 2004). However, further research is required to delineate coincidental improvements in mood and fatigue from true improvements in cognitive function.

### Limitations

One limitation of the current systematic review and meta-analysis relates to lack of assessment of the association between cognitive deficits and depressive

symptoms. The importance of consideration of this association was highlighted in a meta-analysis that revealed that severity of depressive symptoms correlated significantly with impairment across domains of cognition including executive function, episodic memory and processing speed (McDermott & Ebmeier, 2009). However, only a small portion (at most around 10%) of the variability in cognitive function is accounted for by variability in depressive symptom severity (McDermott & Ebmeier, 2009). Therefore, there remains considerable separation between symptoms of depressive mood and cognitive impairment in patients suffering from depression, indicating that cognitive impairment cannot be considered entirely as a secondary feature of low mood in depression. Overall, although there is some evidence of an association between depressive symptomatology and cognitive function, this association does not account for the majority

of variability in cognitive performance in depressed patients.

A further limitation of this study relates to most patients in the included studies being medicated. However, our subanalysis demonstrated significant cognitive deficits in unmedicated currently depressed patients on the SWM, DMS, PRM and RVP tasks, which span the domains of executive function, memory and attention. These findings support the idea that cognitive impairment is at least in part separable from medication effects in currently depressed patients.

The final limitation relates to the range of criteria used to define remission from depression within the remitted samples. Therefore, it is possible that our results may have been affected by the presence of low levels of persisting depressive symptoms in the remitted depressed group.

### Conclusions

This review has demonstrated that cognitive impairment across the domains of executive function and attention, and to an extent memory, represents a core and clinically relevant feature of depression that persists beyond symptoms of low mood. Cognitive impairment is exhibited by depressed patients during current and remitted states, including in unmedicated samples. Previous research has demonstrated that cognitive impairment cannot be fully accounted for by severity of depressive symptoms and, along with symptoms of low mood, is associated with poor psychosocial function. We argue that cognitive impairment may represent a valuable target for new therapies for depression because remediation of cognitive impairment in addition to depressive symptoms will be important in improving functional outcome for patients with depression.

### Supplementary material

For supplementary material accompanying this paper visit <http://dx.doi.org/10.1017/S0033291713002535>.

### Declaration of Interest

Drs Rock, Riedel and Blackwell are full-time employees of Cambridge Cognition, and Dr Blackwell holds shares in Cambridge Cognition. Dr Roiser is a paid consultant for Cambridge Cognition.

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